EXTERNAL SHOCKS AND BANKING CRISES IN DEVELOPING COUNTRIES: DOES THE EXCHANGE RATE REGIME MATTER?

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Abstract

This paper examines some determinants of banking crises in developing economies. Specifically, the effects of terms of trade shocks and capital flows are analyzed. The choice of the nominal exchange rate regime is found to be a crucial factor in the way various shocks are transmitted through the monetary sector. A logit model is used on panel data and preliminary results indicate that countries with flexible regimes were able to lessen the impact of external shocks on the domestic economy. This in turn reduced the likelihood of banking crises.

JEL Classification: E42, E51, G21.

Keywords: banking crises, shocks, exchange rates.

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1. Introduction

The causes and consequences of banking crises have regained prominence after the recent wave of financial and banking crises in emerging economies. A number of internal and external factors, such as capital flows, terms of trade shocks, institutional strength and appreciations of exchange rates have been identified in the literature as contributory factors. While factors such as interest rates, stock market crashes and public confidence could seriously affect the performance of a banking system, the type of exchange rate regime could also be a major determining factor in the way external shocks are transmitted to the banking sector. This is particularly important in small open developing economies, which are heavily dependent on volatile primary product exports and foreign capital and where large negative shocks have the potential to create banking crises.

This paper empirically focuses on the link between external factors and the incidence of banking crises in developing small open economies, (SOEs). Major banking crises over the 1970 to 1992 period are identified from existing case studies. Using theoretical priors from the literature, principal factors that may lead to banking crises in these economies are modeled in a logistic framework. Particular emphasis is placed on the occurrence of terms of trade shocks, capital flows, bank lending and how they affect economies under different nominal exchange rate regimes.

The remainder of the paper is organized as follows. Section II discusses the theoretical literature. Section III analyzes case studies from the literature and the

methodology used to define crisis episodes. Section IV explains the econometric model and Section V conducts the estimation. Section VI concludes the analysis.

II. Literature on Banking Crisis

We begin by defining a banking crisis, with some commonly cited examples:

- "..liquidation of credits that have been built up in a boom." Veblen [1904]
- ".. a sharp reduction in the value of banks' assets, resulting in the apparent or real insolvency of many banks and accompanied by some bank collapse and possibly some runs." Federal Reserve Bank of San Francisco [1985]
- "... situation in which a significant group of financial institutions have liabilities exceeding the market value of their assets, leading to runs and other portfolio shifts, collapse of some financial firms, and government intervention." Sundararajan and Balino [1991]
- ".. a non-linear disruption to financial markets in which adverse selection and moral hazard problems become much worse, so that financial markets are unable to efficiently channel funds to those who have the most productive investment opportunities." Frederic Mishkin [1996]

These definitions imply that banking crises have both micro and macro economic origins. In fact, the interaction of microeconomic and macroeconomic factors could explain a large number of banking crises in SOEs. Macroeconomic factors such as negative terms of trade shocks, level and composition of foreign debt, changes in interest rates, recessions and sudden capital outflows have been suggested as major determinants of crises. Some of these factors are also conditioned by the nature of the policy environment in place, such as institutional strength, confidence in government and the type of the nominal exchange rate regime in operation prior to the occurrence of a crisis. On the microeconomic front,

institutional factors relating to bank supervision and regulation, adequate legal and judicial framework with regard to bankruptcy, law enforcement as well as internationally recognized accounting standards could also have a bearing on the performance and soundness of a banking system. In fact, the combination of several of these factors have the potential to trigger a banking crises.

The theoretical literature for analyzing banking crises is reflective of these numerous contributory factors, with different models and theories explaining various aspects of banking crises. While it is impossible to discuss all models pertinent to banking crises, a brief overview of the main explanations will be conducted. These will be discussed under the monetary approach to financial crises, asymmetric information and micro theories and the business cycle view of banking crises. We start with the monetary view and the role played by exchange rates.

1. Monetary approach and exchange rates

The monetary approach emphasizes the role of money growth and its variability as the principal determinant of a crisis, Friedman and Schwartz [1963]. A financial crisis need not occur at any particular stage of the business cycle, but could result from a change in the monetary base, such as a sudden and erratic tightening of reserve money, or a foreign inflow which may force financial enterprises to sell assets to meet reserve obligations. This may reduce asset prices, raise interest rates and threaten solvency. The exchange rate regime is one of the factors which may affect the way external shocks impact on monetary base and banking sector. This arises works through the demand for money and the supply of supply.

The demand for money is affected through two elements. First there is the conventional change in the transactions demand component. Second, agents also hold money as a store of value. The process by which the supply of money is altered depends on the type of exchange rate regime, i.e. if it is fixed or flexible. If the currency is freely floating, then the supply of money is determined by the central bank. When the exchange rate is fixed, the supply of base money is determined by the balance of payments. The Neary [1985] model analyzes the adjustment process under fixed and floating exchange rates.

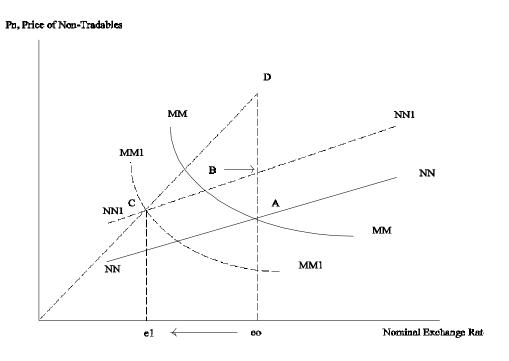
The level of real money balances is determined by a conventional money demand function:

$$(1) m - P = \alpha y - \delta i$$

where i is the domestic interest rate and m, P and y are the logarithms of nominal money demand, the price level and the level of income. This equation is related to the nominal exchange rate in two ways. First, the domestic price level P is a weighted average of the prices of trade and non-traded goods:

$$(2) P = \beta_n P_n + (1 - \beta_n) e.$$

Secondly, the expected changes in the exchange rate influence the link between the domestic interest rate i and the world interest rate i^* (which the home country is too small to



affect). For simplicity, Neary initially ignores expected exchange-rate changes, so domestic and foreign rates are identical. The money market is obtained by substituting the domestic money supply into the money demand function and is depicted in Figure 1. The vertical axis shows the price level of non-tradables. The horizontal axis shows the nominal exchange rate, (i.e. domestic currency per unit of foreign currency). The *NN* locus depicts the non-tradable market equilibrium and *MM* the money market equilibrium. Any ray through the origin corresponds to the relative price of non-traded to traded goods and gives a real exchange rate. The economy's initial equilibrium is at *A*, where the *NN* and *MM* loci intersect.

If the economy operated a flexible exchange rate, the domestic money market is always in equilibrium and the economy must lie along this locus. Under a fixed exchange

rate on the other hand, the economy could be, for example, at a point above the *MM* locus, reflecting a shortfall of actual holdings of real money balances below desired holdings. This disequilibrium must be offset by a buildup of foreign exchange reserves to augment domestic money supply. Therefore, all points above the *MM* locus depict situations of balance of payments surplus and all points below reflect deficits.

The model can be used to analyze the effects of a boom. With a pre-boom equilibrium A, the increased demand for non-tradables shifts the NN locus upwards to NN_I . The increase in real income also raises demand and if domestic money supply is unchanged, the price level must fall to restore money market equilibrium. The liquidity effect shifts the MM locus to MM_I . The nominal exchange rate appreciates from e_0 to e_I , and a new equilibrium C is reached. The greater slope of OC relative to OA implies a fall in the real exchange rate. This combined with the nominal appreciation means that the price of domestic tradable goods unambiguously falls while the price of non-tradables may rise or fall.

Now consider the case when the exchange rate is fixed at e_0 . The price of non-tradables moves the economy to B, and since desired money balances are greater than actual, the equilibrium at this point cannot be maintained indefinitely. Instead the trade surplus leads to a build-up of foreign reserves and in the absence of sterilization, money supply gradually rises. This causes both the MM_1 and NN_1 curves to shift upwards. This process can only end when the post-boom equilibrium real exchange rate is attained at point D where the surplus is eliminated and the economy reaches its new long run equilibrium. In this Neary framework, a fixed exchange rate increases the real and monetary effects of a boom and gives rise to

inflationary pressures as the rise in the price of non-traded goods is brought about by a rise in the nominal price rather than a fall in the price of tradables. If a large part of this monetary expansion is transmitted through the banking system in the form of bank credit, a large debt overhang may result. A slowdown in growth or a recession in subsequent years, may deteriorate the loan portfolio. A negative shock, such as sudden capital outflows would also adversely affect the banking sector. Lower foreign reserves and bank liquidity would result in higher interest rates and subsequently a decline in output and employment. These factors would increase debt servicing burden of borrowers and increase the potential for default. If this is systematic across the financial sector, a banking crisis may result.

The situation may be less acute under a floating regime, if most of the debt is domestic, as capital outflows and reduced demand for real money balances would depreciate the currency and raise domestic prices. This in turn would reduce the real value of assets of the banking system (including loans given to the private sector). Furthermore, the real value of bank liabilities would also fall, lessening the impact of the negative outflow on banks. Floating exchange rates would also accommodate downward wage rigidity through a nominal depreciation, after a negative shock or economic slowdown, easing competitive pressures. The key difference between the fixed and flexible exchange rate scenarios is that the adjustment in the fixed case mainly affects the supply side and the price level while under flexible regimes, the adjustment largely takes place through cahnges in the nominal exchange rate and relative prices.

While the transmission of external shocks through different exchange rate regimes could play a leading role in causing crises, the interaction of various other internal disturbances and institutional factors could also important in determining banking crises.

These relate to institutional strength with regard to supervision, prudential regulation relating to connected lending, accounting standards affecting the disclosure of financial information and an adequate legal environment. Some of these micro theoretic factors are discussed below.

2. Micro theoretic explanations

The micro view comes from asymmetric information and credit market analysis. The most commonly discussed approach is the credit rationing situation resulting from various market failures, Stiglitz and Weiss [1981]. Due to asymmetric information and adverse selection, banks may ration credit, creating problems for non-financial firms. According to Mishkin [1996], the problems of moral hazard and adverse selection rise after stock market crashes. As the value of net worth declines, the moral hazard problem increases as borrowers have less to lose by making a more risky investment. Demirguc-Kunt and Detragiache, hereafter DKD [1997], have argued that these problems of moral hazard and adverse selection are attenuated after financial liberalization in developing countries. While the benefits of financial liberalization have been well documented in theoretical and empirical work, hasty liberalization of often weak financial sectors have known to create financial crises in subsequent years, Diaz-Alexandro [1985]. DKD [1997] address this issue using cross country data and shows that countries which liberalized with weak institutional and regulatory frameworks were more susceptible to financial crises in subsequent years.

3. Business cycle explanations of banking crises

The business cycle approach looks at the vulnerability of a financial sector over the business cycle. The financial sector responds endogenously to movements in the business cycle, see Minsky [1977], Taylor and O'Connell [1985]. A crisis may develop due to systematic forces near the peak of the cycle as interest rates rise. Reduced lending by banks and high interest payments may adversely affect non-financial firms, increasing the likelihood of default. The entire financial sector may suffer after a surprise shock, such as news of a major bankruptcy. Studies by Calomiris and Garton [1991], Greenwald and Stiglitz [1988] and Bernanke and Gertler [1988] show that unanticipated shocks such as stock market crashes could lead to financial crises through the impact on the balance sheets of financial institutions.

External conditions could also affect assets of the banking sector. Terms of trade shocks could profoundly affect the profitability of firms and households in primary product exporting economies. Unanticipated changes in the terms of trade could make debtors unable to discharge their debts, deteriorating bank balance sheets. Sachs, Tornell and Velasco [1996] argue that countries with a high proportion of short term debt may end up with a maturity mismatch due to sudden changes in interest rates and debt service requirements. Eichengreen and Rose [1998] make the point that this maturity mismatch is attenuated in developing countries where the average length of maturity is much shorter than in developed countries. Aligned to this is the problem of currency mismatch, where a large proportion of loans is denominated in foreign currency, as an exchange rate depreciation may severely increase the debt service requirement.

III: Evidence of Banking Crises

The literature has mainly focused on developed economies, particularly the United States, to analyze numerous episodes of bank runs and crises, with the most notable during the Great Depression, Benanke [1983], Hubbard [1991]. There have been severe banking crises in other developed countries, Honohan [1997]. Developing countries have also been subject to major crises over the last 20 years, with considerable costs attached, Caprio and Klingebiel [1996], (Table 1). Due to the small and concentrated nature of banking systems in developing countries, even a single bank failure can have tremendous effects.

Table 1
Developing Country Banking Crises - Post 1970

| Country | Extent of banks affected |
|---------------|--|
| Cameroon | High proportion of loans written off |
| Chile | 7 commercial banks, 1 finance company affected |
| Colombia | 6 banks affected |
| Cote D'Ivoire | 4 big banks insolvent |
| Ecuador | one bank liquidated, one take over |
| Egypt | 5 banks affected |
| Gabon | state banks affected |
| Ghana | 7 banks insolvent |
| Honduras | 12 banks affected |
| Jordan | large number of banks affected |
| Kenya | 4 banks, 24 non-bank firms in distress |
| Malawi | Lending to agricultural parastatsals |
| Malaysia | 4 banks insolvent, 24 others affected |
| Mauritania | S major banks |
| Mexico | 29 banks closed or merged |
| Nigeria | half banks in distress |
| Philippines | 8 banks, 32 thrifts, 128 small banks |
| Senegal | 6 commercial banks and 1 development bank |
| Sri Lanka | 4 state owned banks |
| Tanzania | most of banking system insolvent |
| Thailand | 24 finance companies closed, 3 commercial banks affected |
| Uganda | 50% of banks in distress |
| Uruguay | numerous banks affected |
| Venezuela | many banks were affected, including US branches |
| Zambia | Meridian Bank became insolvent |

Source: Caprio and Klingebiel [1996], Sundararajen and Balino [1991]

Table 2
Nature of Developing Country Banking Crises

| Country | Cause(s) | Costs of restructuring |
|---------------|--|----------------------------------|
| Cameroon | oil shock | |
| Chile | high RER & interest rates, tot shock, | 41.2%% of GDP |
| | asset bubble, wage rigidity | 45% of financial assets |
| Colombia | bad lending & management, low growth | 25% of assets, 5% of GDP |
| | high debt, poor regulations k institutions | |
| Cote d'Ivoire | tot shock, inadequate | 60-70% of banking assets |
| | supervision, high RER | 25% of GDP |
| Ecuador | recession, high debt levels and interest rates | na |
| Egypt | oil shock | 4.9% of GDP |
| Gabon | oil shock | 5.6% of GDP |
| Ghana | poor regulations, supervision, devaluation | 6% of GDP |
| Honduras | inadequate monitoring | na |
| Jordan | fall in private capital inflows | na |
| Kenya | tot shocks, drought, connected lending, | 15% of liabilities |
| | insufficient capitalized banks, rules | |
| Malawi | poor regulations and lending | na |
| Malaysia | tot shocks, fall in asset prices | 7.7% of deposits, 4.7% of GDP |
| Mauritania | tot shock | na |
| Mexico | excessive borrowing, high interest rates | 2.3% of GDP |
| Nigeria | poor regulations, debt overhand | 20% of bank assets |
| Philippines | tot shocks, bad lending | 5.2% of deposits, 3% of GDP |
| Senegal | tot shock, drought, | 20-30% of financial assets |
| | poor supervision | 17% of GDP |
| Sri Lanka | tot shocks, poor lending and regulation | 35% of loan portfolio, 5% of GDP |
| Tanzania | excessive parastatal lending | 10% of GDP |
| Thailand | poor management & regulations | 25% of assets |
| | oil shock, | 0.7% of GDP |
| Uganda | bad lending | na |
| Uruguay | fall in beef prices, high interest rates, | na |
| Venezuela | financial liberalization, poor loans, | 13% of GDP, 30% of deposits |
| Zambia | poor management and regulations | 13% of commercial bank assets |
| | | cost - 1.4% of GDP |

Sources: Caprio and Klingebiel [1996] na- data not given in case studies

Table 3 describes how crisis years for each country were assigned. The first category consists of studies where crisis years were explicitly stated, namely by Caprio and Klingebeil [1996] and DKD [1997]. They used one of the following four main thresholds to define a

banking crisis; (i) that the ratio of non-performing loans to GDP must exceed 10%, or (ii) the cost of the rescue operation must be at least 2% of GDP, or (iii) there must be a large scale reorganization and nationalization of banks, or (iv) the enactment of various emergency measures, such as deposit freezes, prolonged bank holidays, deposit guarantees, etc.

Table 3
Definition of Banking Crises

| Country | Crisis Years I | Explicitly Stated | Crisis years im | plicitly ascertain | ed from cas | se studies | |
|---------------|----------------|-------------------|-----------------|--------------------|-------------|--------------------|--------------|
| | Caprio | Demirgu-Kunt | Sundararajan | Hausmann & | Morris et | Alexander | Crisis years |
| | & Klingebeil | & Detragiache | & Balino | Rojas-Suarez | al_1 | et al ₂ | used |
| Cameroon | 1987-90 | | | | | 1986-90 | 1986-90 |
| Chile | 1981-83 | | 1981-84 | | | | 1981-84 |
| Colombia | 1982-87 | | | 1982-85 | | | 1982-85 |
| Cote D'Ivoire | 1988-91 | | | | | 1987-92 | 1987-92 |
| Ecuador | early 1980s | | | | 1984-88 | | 1984-88 |
| Egypt | 1987-92 | | | | | 1987-92 | 1987-92 |
| Gabon | | | | | | 1986-90 | 1986-90 |
| Ghana | 1982-89 | | | | | 1982-90 | 1983-90 |
| Honduras | | | | | 1983-92 | | 1983-92 |
| Jordan | | | | | | 1989-90 | 1989-90 |
| Kenya | 1985-89 | | | | | | 1986-89 |
| Malawi | | | | | | 1981-85 | 1981-85 |
| Malaysia | 1985-88 | 1985-88 | | | | 1986-88 | 1985-88 |
| Mauritania | 1984-93 | | | | | | 1988-92 |
| Mexico | 1981-82 | 1982 | | | | | 1982 |
| Nigeria | 1990s | 1991-94 | | | | | 1990-92 |
| Philippines | 1981-87 | 1981-87 | 1981-87 | | | | 1981-87 |
| Senegal | 1988-91 | 1983-88 | | | | | 1984-88 |
| Sri Lanka | 1989-93 | 1990-93 | | | | | 1990-92 |
| Tanzania | 1987 | 1988-94 | | | | | 1988-92 |
| Thailand | 1983-87 | | 1983-87 | | | 1983-87 | 1983-87 |
| Uganda | 1994 | 1990-94 | | | | | 1990-92 |
| Uruguay | 1981-84 | 1981-85 | 1981-86 | | | | 1981-86 |
| Venezuela | 1980 | 1993-94 | | | | | 1980-81 |
| Zambia | 1991-95 | | | | | | 1991-92 |

The second group of studies analyzed banking crisis years implicitly, with qualitative information relating to financial fragility, non-performing loans and the cost of restructuring. Although the information was not comprehensive, these studies gave an indication of the

severity of crises. For the purposes of this paper, common crisis periods were identified from all six studies. When the crisis periods differed, information from the studies discussing the crises implicitly were used as they had more detailed in-depth information.

IV: Empirical Estimation

1. Econometric model

The econometric model estimates the probability of a country experiencing a crisis using a logit model. A logit model estimates whether or not an event occurs, or in this case whether a country experienced a crisis or not. Following Baltagi [1995], the dependent variable is a binary choice variable $y_{it} = 1$ if the event happens and 0 if it does not happen for country i at time t. If p_{it} is the probability that a crisis occurs, then $E(y_{it}) - 1.p_{it} + 0.(1-p_{it}) = p_{it}$. This is usually modeled as a function of some explanatory variables:

(3)
$$p_{it} = \Pr[y_{it} = 1] = E(y_{it} \mid x_{it}) = F'(x'_{it} \mid \beta).$$

For the linear probability model, $F(x'_{it}\beta) = x'_{it}\beta$, the usual panel data methods apply except that \hat{y}_{it} is not guaranteed to lie in the unit interval. The standard solution has been to use the logistic or normal cumulative distribution functions that constrain $F(x'_{it}\beta)$ to lie between 0 and 1. In this case, a country experiences a crisis if the explanatory variable(s) exceeds some unobserved threshold, i.e.

13

(4)
$$y_{it} = 1$$
, if $y*_{it} > 0$,

(5)
$$y_{it} = 0$$
, if $y^*_{it} \le 0$,

where $y_{it} = x_{it}^{\prime} \beta + u_{it}$, so that

(6)
$$\Pr[y_{it} = 1] = \Pr[y_{it}^* > 0] = \Pr[u_{it} > -x_{it}^* \beta] = F(x_t^* \beta).$$

The last equality holds as long as the density function describing F is symmetric around zero.

2. Fixed effects estimation

Moving from the pooled estimation, a useful extension of the analysis is the fixed effects estimation with country characteristics. Chamberlin [1980] suggests a way of wiping out deviations from group means in a logit framework. Consider a large sample with *n* observations, and *T* time periods. Chamberlin [1980] suggests using the following conditional likelihood function to get a computationally convenient estimator:

(7)
$$L = \prod_{i=1}^{N} \Pr(y_{i1}, \dots, y_{iT} \mid \sum_{t=1}^{N} y_{it}).$$

This implies that the likelihood for each set of T observations is conditioned on the number of Is in the panel. By conditioning on the sum of observations, heterogeneity effects can be removed and a conditional likelihood function created from the product of those terms for which the sum is not zero or T. For example, let us consider the case where T=2; the unconditional likelihood is

(8)
$$L = \prod_{i=1}^{N} \Pr(y_{i1}) \Pr(y_{i2}).$$

The sum $(y_{i1} + y_{i2})$ can be 0, 1 or 2. If it is 0, both y_{i1} and y_{i2} are 0 and

(9)
$$\Pr[y_{i1} = 0, y_{i2} = 0 \mid y_{i1} + y_{i2} = 0] = 1.$$

Similarly, if the sum of both y_{i1} and y_{i2} are 1 and

(10)
$$\Pr[y_{i1} = 1, y_{i2} = 1 \mid y_{i1} + y_{i2} = 2 = 1].$$

Since log 1 =0, these terms add nothing to the conditional likelihood. Only observations for which $y_{i1} + y_{i2} = 1$ matter in log L are given by

(11)
$$\Pr[y_{i1} = 0, y_{i2} = 1 \mid y_{i1} + y_{i2} = 1],$$

and

(12)
$$\Pr[y_{i1} = 1, y_{i2} = 0 \mid y_{i1} + y_{i2} = 1].$$

The latter can be calculated as

(13)
$$\Pr[y_{i1} = 1, y_{i2} = 0 \mid \Pr[y_{i1} + y_{i2} = 1],$$

with

(14)
$$\Pr[y_{i1} + y_{i2} = 1] = \Pr[y_{i1} = 0, y_{i2} = 1] + \Pr[y_{i1} = 1, y_{i2} = 0],$$

since the latter two events are mutually exclusive. Therefore,

(15)
$$\Pr[y_{i1} = 1] = \frac{e^{u_i + x_{ii}'\beta}}{1 + e^{u_i + x_{ii}'\beta}}.$$

This means that,

(16)
$$\Pr[y_{i1} = 1, y_{i2} = 1]y_{i1} + y_{i2} = 1] = \frac{e^{x_{i1}\beta}}{e^{x_{i2}\beta}}.$$

Similarly,

(17)
$$\Pr[y_{i1} = 0, y_{i2} = 1 \mid y_{i1} + y_{i2} = 1] = \frac{e^{x'_{i2}\beta}}{e^{x'_{i2}\beta} + e^{x'_{i2}\beta}},$$

and neither probability involves u. By conditioning on $y_{il} + y_{i2}$, the u_i have been swept away. We now move to the estimation.

3. The variables

Definition of variables and the corresponding descriptive statistics are shown in Tables 4 and 5.

Table 4

| | Description of variables | |
|---------------|---|-------------------------|
| Varaible Name | Description | Source |
| GDPGR | Real GDP growth in logs | IFS: line 99b |
| Trade Shock | Percentage shock measure | World Tables, UNTACD |
| TOT Trend | Terms of trade trend | World Tables, UNTACD |
| RMOGR | Real growth in MO | IFS: line 14 |
| RM1GR | Real growth in MI | IFS: line 34 |
| RM2GR | Real growth in M2 | IFS: line 35 |
| DEFGR | Growth in GDP deflator | IFS: line 99bip |
| FISGDP | Ratio of fiscal deficit to GDP | IFS: line 80 |
| RESGDP | Ratio of foreign reserves to GDP | IFS: line 79dad |
| INFLATION | Inflation rate proxied by the annual CPI | IFS: line 64 |
| CAPFLOW | Ratio of foreign capital inflows | World Debt Tables |
| | [aggregate] to GDP | |
| M2RES | Ratio of M2 money to international reserves | IFS: lines 35179dad |
| LEND | Ratio of private sector credit to GDP | IFS: lines 99b122d |
| DEBTGDP | Ratio of external total debt to GDP | Frankel and Rose (1997) |
| AVINTDEBT | Average of interest rates faced by country | Frankel and Rose (1997) |
| | on foreign debt | |
| LONGDEBT | Ratio of foreign long term debt to total debt | Frankel and Rose (1997) |
| COMDEBT | Ratio of foreign commercial bank debt | Frankel and Rose (1997) |
| | to total debt | |
| CONDEBT | Ratio of concessional debt to total debt | Frankel and Rose (1997) |
| VARDEBT | Ratio of foreign variable debt to total debt | Frankel and Rose (1997) |
| LENINFRA | Interactive term between infrastructure | BERI |
| | Index and lending | |
| LENBUR | Interactive term between bureaucratic | ICRG |
| | delay index and lending | |
| CRISES | [1,01 dummy variables for banking crises | Various case studies |

Table 5
Descriptive Statistics

| Variable | Obs. | Mean | St. Dev. | Max. | Min |
|-----------|------|-------|----------|-------|--------|
| | | | | | |
| GDPGR | 899 | 0.01 | 0.06 | 20 | 19.98 |
| SHOCK | 895 | 2.62 | 20.85 | -60.2 | 33.3 |
| RM0GR | 1035 | 0.06 | 0.26 | 3.77 | -1.96 |
| RM1GR | 1059 | 0.01 | 0.3 | 5.98 | -1.89 |
| RM2GR | 1058 | 0.1 | 0.3 | 5.37 | -1.3 |
| FISGDP | 975 | -0.01 | 0.07 | 0.62 | -0.25 |
| RESGDP | 807 | 0.01 | 0.04 | 0.26 | -0.5 |
| DEFGR | 1105 | 0.11 | 0.23 | 4.73 | -1.94 |
| INFLATION | 1055 | 0.19 | 0.36 | 5 | -0.13 |
| CAPFLOW | 880 | 0.08 | 0.09 | 0.6 | -0.061 |
| M2RES | 1024 | 4.71 | 26.46 | 812.7 | 0.04 |
| LEND | 917 | 0.24 | 0.18 | 0.9 | 0.002 |
| DEBTGDP | 886 | 0.83 | 1.52 | 13.9 | 0.01 |
| AVINTDEBT | 1006 | 6.07 | 2.82 | 16.5 | 0.12 |
| LONGDEBT | 858 | 0.49 | 0.46 | 4.2 | 0.02 |
| COMDEBT | 858 | 0.11 | 0.15 | 0.94 | 0.01 |
| VARDEBT | 858 | 0.13 | 0.18 | 1.1 | 0.01 |
| CONDEBT | 858 | 0.19 | 0.28 | 2.08 | 0.02 |
| LENINFRA | 285 | 0.61 | 0.56 | 2.94 | 0.003 |
| LENBUR | 285 | 0.7 | 0.58 | 2.94 | 0.001 |
| | | | | | |

3.11. Measuring external shocks

The impact of shocks is captured by the terms of trade and capital inflows.

Unanticipated terms of trade shocks are measured by the deviation of the terms of trade from its long run trend and is expressed as a percentage term. This variable is further disaggregated into positive and negative shocks in subsequent sensitivity analysis. The "size" of these shocks is therefore regressed against the crisis variable. The impact of capital inflows is captured by the capflow variable, the ratio of capital flows to GDP. Capital flows consist of net long term debt flows (commercial and public), net foreign direct investment flows and portfolio flows excluding official aid.

Table 6
Identifying Shock Periods

| Country | Terms of trade | | Exchange rate regir | |
|---------------|----------------|-----------|---------------------|----------------|
| | Positive | negative | positive shock | negative shock |
| Botswana | 1977-1982 | | | Flexible |
| Cameroon | 1978-1982 | 1986-1992 | Fixed | Fixed |
| Chile | 1968-1974 | 1975-1988 | Fixed | Fixed |
| Colombia | 1976-79 | | Flexible | |
| Congo | 1972-1975 | 1986-1990 | Fixed | Fixed |
| | 1979-1985 | | Fixed | |
| Costa Rica | 1976-1979 | | Fixed | |
| Cote D'Ivoire | 1976-1980 | | Fixed | |
| Dominican Rep | 1974-76 | | Fixed | |
| Ecuador | 1977-80 | 1986-90 | Fixed | Fixed |
| | 1982-86 | | | |
| Egypt | 1972-74 | | Fixed | |
| | 1979-85 | 1986-90 | Fixed | Fixed |
| El Salvador | 1977-80 | | Fixed | |
| Ethiopia | 1976-80 | | Fixed | |
| Gabon | 1979-84 | 1986-92 | Fixed | Fixed |
| Ghana | 1976-79 | | Fixed | |
| Guatemala | 1976-1980 | | Fixed | |
| Indon | 1979-85 | 1986-90 | Flexible | Flexible |
| Korea | | | | |
| Kenya | 1976-1980 | | Fixed | |
| Malawi | 1976-1980 | 1981-83 | Fixed | Fixed |
| Malaysia | 1977-1985 | 1986-90 | Fixed | Fixed |
| Mauritius | 1974-1977 | | Fixed | |
| Mexico | 1979-85 | 1986-90 | Flexible | |
| Morocco | 1974-78 | | Fixed | |
| Niger | 1970-74 | | Fixed | |
| Nigeria | 1972-75 | | Flexible | |
| | 1979-85 | 1986-90 | Flexible | Flexible |
| Paraguay | | | Fixed | |
| Philippines | 1973-75 | 1979-85 | Flexible | |
| Senegal | 1974-1979 | 1979-84 | Fixed | Fixed |
| Sri Lanka | 1976-79 | | Flexible | |
| Syria | | | Fixed | |
| Tanzania | 1976-80 | | Fixed | |
| Thailand | 1974-76 | 1980-85 | Fixed | |
| Tunisia | 1974-78 | 1981-85 | Fixed | |
| Uganda | 1976-79 | | Fixed | |
| Uruguay | 1973-75 | 1980-85 | Flexible | |
| Venezuela | 1973-1977 | | Fixed | |
| | 1979-85 | 1986-90 | Fixed | |
| Zambia | 1969-72 | | Fixed | |
| | 1974-76 | 1977-85 | Fixed | |

3.12. Financial variables

With the lifting of controls on interest rates and directed credit, evidence from case studies points to lending booms with the proliferation of new banks. A corollary of this may be a worsening of financial fragility, especially if undertaken without adequate institutional development, DKD [1997]. The literature uses a number of variables to proxy for financial liberalization. An obvious choice is lending itself (Lend), which is private sector credit from the banking sector expressed as a percentage of GDP. To test whether sudden capital outflows lead to liquidity crises, the ratio of M2 money to foreign reserves is introduced, (M2reserves). The composition of debt, in terms of maturity and creditor has received much attention, Sachs et al [1996]. This may be important for countries with a high proportion of short term foreign commercial debt with variable interest rates.

3.13. Macroeconomic variables

Inflation was introduced as it is normally associated with mismanagement of the economy. Other macroeconomic variables include the fiscal deficit and various measures of nominal exchange rate in the sensitivity analysis. The real exchange rate (Reer) was used to test for overvaluation. Real GDP growth was used to investigate the impact of banking crises on real income . In order to establish the causality between real income growth and crises, lagged values of the real GDP variable were used in the predictive model. Finally, to test for sensitivity of crisis year definitions, different years of banking crises were introduced.

V. Econometric Results

1. Baseline model

Table 7 shows the results from the baseline case, using Stata (5).

Table 7 Logit regressions

| | Baseline | | External s | hocks | Instituti | ons |
|-------------------------|----------------------|-------|---------------------|-------|--------------------|-------|
| Dep. var: Crises | $\delta F(x)/x$ | z | $\delta F(x)/x$ | z | $\delta F(x)/x$ | z |
| Constant | 6.527 | 2.737 | 8.315 | 3.418 | -4.928 | 1.092 |
| Macro variables | | | | | | |
| Rgdpgr | -4.732 | 2.290 | -5.055 | 2.372 | -12.544 | 2.608 |
| Inflation | -0.076 | 0.153 | -0.113 | 0.219 | -3.419 | 2.228 |
| Reer | -0.005 | 2.240 | -0.005 | 2.272 | 0.008 | 0.547 |
| Financial variables | | | | | | |
| Lend | 0.608 | 0.809 | 0.848 | 1.103 | 2.753 | 0.557 |
| M2res | 0.028 | 2.54 | 0.025 | 2.294 | -0.912 | 0.428 |
| Debtgdp | 0.492 | 3.496 | 0.699 | 5.091 | 2.167 | 2.937 |
| Avintdebt | 0.113 | 2.234 | 0.100 | 1.985 | 0.076 | 0.846 |
| Shock variables | | | | | | |
| Tot trend | -1.974 | 3.785 | -2.368 | 4.324 | 0.343 | 0.356 |
| Trade shock | 0.089 | 1.096 | | | -0.035 | 2.042 |
| Capflow | 2.080 | 1.357 | | | 0.241 | 0.051 |
| Floatshock | | | -0.088 | 3.838 | | |
| Intershock | | | 0.036 | 2.372 | | |
| Fixedshock | | | 0.018 | 2.020 | | |
| Floatcap | | | -19.976 | 1.941 | | |
| Intercap | | | 8.456 | 1.289 | | |
| Fixedcap | | | -2.991 | 0.945 | | |
| Lenbure | | | | | -7.065 | 2.481 |
| Leninfra | | | | | 5.604 | 1.790 |
| Pseudo R squared | 0.147 | | 0.1934 | | 0.227 | |
| No of observations | 674 | | 674 | | 222 | |
| | | Prob | | Prob | | Prob |
| H0: Slopes =0 | $\chi^2(10) = 78.59$ | 0.00 | χ^2 (14)=76.84 | 0.00 | $\chi^2(12)=50.1$ | 0.00 |
| H0: Macro effects=0 | χ^2 (3)110 | 0.01 | χ^2 (3)= 11.82 | 0.01 | χ^2 (3)=10.82 | 0.01 |
| H0: Financial effects=0 | χ^2 (4)= 26.51 | 0.00 | χ^2 (4)= 35.94 | 0.00 | χ^2 (5)=7.65 | 0.17 |
| H0: Shock effects $=0$ | χ^2 (3)=18.55 | 0.00 | χ^2 (7)= 38.41 | 0.00 | χ^2 (3)=5.92 | 0.12 |

| Goodness of fit model | S | (Cut off probability | probability 0.5) | | |
|-----------------------|---------------------|----------------------|------------------|-------|--|
| | | Crisis | No Crisis | Total | |
| Baseline model | Predicted Crisis | 12 | 10 | 22 | |
| | Predicted no Crisis | 79 | 573 | 652 | |
| | Total | 91 | 583 | 674 | |
| Shocks model | Predicted crisis | 16 | 13 | 29 | |
| | Predicted no crisis | 75 | 576 | 645 | |
| | Total | 91 | 583 | 674 | |
| Institutions model | Predicted crisis | 14 | 7 | 21 | |
| | Predicted no crisis | 31 | 169 | 200 | |
| | Total | 45 | 176 | 221 | |

The coefficients and associated z-statistics are shown in the first two columns. Diagnostic tests are shown at the bottom of the table together with joint hypothesis tests for the shock, macro and financial effects. Finally, tabulations of actual and predicted values are reported for each regression. Overall, the pooled results have a low explanatory power. However, the levels are similar to those found in the literature, for example Eichengreen and Rose [1998]. Of the macro variables, the contemporaneous growth rate is highly significant, implying that negative real income shocks could lead to a banking crisis. However, as causality may run from the banking sector to the real economy, a predictive model with lagged values is tested in the next section.

Coming to the other macro variables, the inflation term was not significant, while the real exchange rate (Reer) term was significant and negatively associated with crises. This implies that real exchange rate appreciations increases the probability of banking crises. The ratio of foreign debt to GDP was highly significant. Furthermore, the average interest rate faced by a country on its foreign debt (Avindebtgdp), was also significant. Sudden increases in debt service requirements strongly increased the probability of banking crises, supporting Sachs et al. [1996], that debt maturity mismatches could generate banking crises. Although DKD [1997] also find real interest rates to be significant, the interest rate used in this study applies to the debt stock outstanding for that particular country, as changes in this variable are a more appropriate indicator of a crisis than a general interest rate. The trade shock variable was positive, but insignificant, though the trend term was significant. Capital inflows were also not significant. The other significant variable is the ratio of M2 money to international reserves. The results suggest that the probability of a crisis is significantly

enhanced with a low level of reserves, i.e. a sudden capital outflow could seriously undermine the banking system. Looking at the diagnostic statistics, the test for joint significance for the aggregate effects was jointly significant in the baseline model. From the table of actual and predicted outcomes, at a cut-off probability of 0.5%, 12 out of 22 cases were correctly predicted as having crises and 573 out of 652 cases were correctly predicted as not having crises in the baseline model, i.e. in total 86% of the cases were correctly classified.

2. Shocks and choice of exchange regime

The relationships between nominal exchange rate regimes and shocks was analyzed to investigate the nature of the monetary transmission channel. Nominal regimes were broadly classified into fixed pegs, intermediate regimes and floating rates. Dummy variables were created for each year for the three types of regime. These in turn were interacted with the shock variables. For example, if Kenya experienced a positive external shock in 1979, and the exchange rate regime in operation was a fixed peg, then the interactive dummy for the fixed exchange rate and the external trade shock, (Fixshock), took a value of one multiplied by the shock variables, and zero for other years. This differs from the Eichengreen and Rose study, which employs [1,0] dummy variables for each regime. The question posed in this study is more specific, investigating how different regimes interact with external shocks.

The results in the second regression in Table 7 also support the priors on the effects of trade shocks on the macro economy. Terms of trade shocks that enter through floating exchange rate regimes decrease the probability of banking crises. For both intermediate and pegged regimes, the results are diametrically opposite, where shocks significantly increase the probability of crises. The results are less conclusive for capital inflows, though the interactive variable on the floating exchange rate term is negative and significant. An interpretation could be that capital flows entering through floating exchange rate regimes are less likely to cause crises. These results again support the assertions made in the theory regarding the transmission of external inflows and their monetary consequences. It was suggested that under fixed exchange rate regimes, external inflows would have a greater effect on monetary growth, particularly on the supply side. Shocks and capital inflows going though floating regimes had a lower probability of creating a crisis in subsequent years (with a negative sign), than shocks that went through more rigid regimes, which had a positive sign on the coefficients. Coming to the diagnostics, the combined macro, shock and financial effects were jointly significant. In the goodness-of-fit table, 576 out of the 645 cases were correctly predicted as not having a crisis, and 16 out of the 29 cases were predicted as having crises, i.e. (87%) of the case were correctly classified.

3. Role of institutions

Since theoretical foundations concerning asymmetric information and moral hazard are linked to institutional structure, the level of "financial institutional development" could significantly affect banking crises. Unfortunately, data on institutional development, such as connected lending, corruption, bureaucracy, prudential regulations and supervision, are non-

existent across countries. The closest proxies available were the Bureaucratic Delay Index from the Business Environmental Risk Intelligence (BERI) organization, and the International Country Risk Guide (ICRG) measures on infrastructure.

These variables consist of indices specifying infrastructure quality, on a 0 to 4 scale. Higher values imply low infrastructure quality. However, complete data for all the countries were not available, with data missing for most of the 1970s and for the entire period for some countries. This reduced the sample size to 222 observations. The infrastructure index was interacted with the lending variable to yield (Leninfra), to proxy how poor levels of institutions such as prudential regulation and banking supervision might have led to excessive lending. The signs on the coefficient imply that the level of infrastructure development could affect the probability of banking crises through bank lending, i.e. lending booms associated with low quality institutions increased the likelihood of crises. Interacting the bureaucratic variable with lending, (Lenbur), did not yield significant results. The signs on the other key variables such as real income, and shock variables were significant for the institutions equation. The other significant variable is the lagged ratio of M2 money to international reserves. DKD [1997] also use an institutional variable in their study, where a law and order variable measuring the quality of law enforcement, i.e. measures of effective legal and judiciary systems, (proxying corruption) was highly significant, i.e. a higher value in the index implies a higher level of law and order which decreases the probability of crises.

From the table of predictions, 14 out of the 21 cases were correctly predicted as having a crisis, while 169 out of the 200 cases were correctly predicted as not having crises,

implying in total that 83% of all the cases were correctly called. However, these results must be treated with caution due to the small sample size. The models were then subjected to a range of robustness and sensitivity tests.

4. Predictive model

When lagged values of the explanatory variables were used, they more or less confirmed the results of the baseline model. The average rate of interest variable, debt to GDP ratio and M2 to GDP were all correctly signed and significant. Eichengreen and Rose

Table 8 Predictive models

| | Baseli | | Externa | ıl shocks |
|-------------------------|----------------------|-------|---------------------|-----------|
| Dep. var: Crises | $\delta F(x)/x$ | z | $\delta F(x)/x$ | z |
| Constant | 5.201 | 2.148 | 6.147 | 2.324 |
| Macro variables | | | | |
| Lagdpgr | -4.689 | 2.227 | -5.453 | 2.441 |
| Lainflation | -0.143 | 0.292 | -0.529 | 0.855 |
| Lareer | -0.003 | 1.681 | -0.003 | 1.524 |
| Financial variables | | | | |
| Lalend | 0.139 | 0.175 | 0.137 | 0.157 |
| LaM2res | 0.027 | 2.423 | 0.030 | 2.567 |
| Ladebtgdp | 0.423 | 2.959 | 0.504 | 3.045 |
| Lavintdebt | 0.158 | 3.098 | 0.161 | 2.920 |
| Shock variables | | | | |
| Latot trend | -1.777 | 3.367 | -2.01 | 3.496 |
| Latrade shock | 0.003 | 0.357 | | |
| Lacapflow | 3.560 | 2.192 | | |
| Lafloatshock | | | -0.134 | 4.093 |
| Laintershock | | | 0.014 | 0.749 |
| Lafixedshock | | | 0.014 | 1.522 |
| Lafloatcap | | | -6.645 | 1.070 |
| Laintercap | | | 14.945 | 4.135 |
| Lafixedcap | | | 2.796 | 1,379 |
| Pseudo R squared | 0.1363 | | 0.2276 | |
| No of observations | 634 | | 634 | |
| | | Prob | | Prob |
| HO: Slopes = 0 | χ^2 (10)= 56.93 | 0.00 | χ^2 (14)=73.61 | 0.00 |
| HO: Macro effects 0 | χ^2 (3)=8.31 | 0.04 | $\chi^2(3) = 9.17$ | 0.03 |
| HO: Financial effects 0 | χ^2 (4)= 25.91 | 0.00 | χ^2 (4)= 26.03 | 0.00 |
| HO: Shock effects $= 0$ | χ^2 (3)=19.79 | 0.00 | χ^2 (7)= 45.81 | 0.00 |

| Goodness of fit mode | ls | (Cut off probability | 0.5) | |
|----------------------|---------------------|----------------------|-----------|-------|
| | | Crisis | No Crisis | Total |
| Predictive model | Predicted Crisis | 11 | 12 | 23 |
| | Predicted no Crisis | 78 | 533 | 611 |
| | Total | 89 | 545 | 634 |
| Shocks model | Predicted crisis | 23 | 12 | 35 |
| | Predicted no crisis | 66 | 533 | 599 |
| | Total | 89 | 545 | 634 |

[1998] also find high debt to GDP ratios and high foreign interest rates to be significant predictors of crises. The lagged value of capital flows is a significant predictor of impending problems. In the disaggregated shocks model, the effects of shocks transmitted through floating exchange rate regimes reduced the likelihood of a crisis. While the lagged floatcap term became insignificant, the value of the intermediate interactive term became highly significant. Coming to the test of joint significance, only the macro effects were rejected at the 3% and 4% levels. Furthermore, 11 out of 23 outcomes were correctly predicted as having crises and 533 out of 611 were predicted as not having crises, yielding a combined 84% percentage of correctly predicted outcomes for the baseline variant. By the same token, 88% of cases were correctly predicted in the disaggregated model.

5. Robustness

Regional dummy variables were introduced to test for robustness. The results more or less remain unchanged as shown by the Table 9. Both regional dummies were insignificant. The tests for joint significance also suggest robust results except for the institutions regressions, where the joint test for shock effects is rejected at the 5 and 10 per cent levels. Finally, from the goodness-of-fit tables, high levels of predicted outcomes were obtained.

The coefficient on the real income term remained significant. However, the real exchange rate becomes significant in the second equation, suggesting that the causality in adjustment runs from the crisis to the real exchange rate in this version of the analysis. Both the ratio of M2 to reserves and debt to GDP become significant "during" the crisis. The average interest rate faced by countries on the other hand loses significance. Looking at the specification of the model, the joint significance tests are rejected at the 5% level suggesting that the second model gives a better fit.

Table 9 Robustness

| | Baseline | model | External s | hocks | Institu | tions |
|---------------------|-----------------|-------|-----------------|-------|-----------------|-------|
| Dep. var: Crises | $\delta F(x)/x$ | z | $\delta F(x)/x$ | ΙzΙ | $\delta F(x)/x$ | z |
| Constant | 7.523 | 3.069 | 9.307 | 3.662 | -5.62 | 1.201 |
| Macro variables | | | | | | |
| Rgdpgr | -5.371 | 2.515 | -5.496 | 2.519 | -12.964 | 2.620 |
| Inflation | -0.116 | 0.225 | -0.133 | 0.255 | -3.377 | 2.134 |
| Reer | -0.006 | 2.508 | -0.006 | 2.479 | 2 | 0.676 |
| Financial variables | | | | | | |
| Lend | 0.203 | 0.252 | 0.625 | 0.768 | 4.189 | 0.785 |
| M2res | 0.034 | 2.894 | 0.029 | 2.507 | -0.037 | 0.114 |
| Debtgdp | 0.465 | 3.180 | 0.684 | 4.935 | 2.182 | 2.859 |
| Avintdebt | 0.110 | 2.508 | 0.089 | 1.634 | 0.079 | 0.740 |
| Shock variables | | | | | | |
| Tot trend | -2.03 | 3.840 | -2.44 | 4.422 | 0.543 | 0.536 |
| Trade shock | 0.009 | 1.111 | | | -0.369 | 2.164 |
| Capflow | 2.763 | 1.716 | | | 1.372 | 0.247 |
| Floatshock | | | -0.084 | 3.606 | | |
| Intershock | | | 0.035 | 2.364 | | |
| Fixedshock | | | 0.019 | 2.046 | | |
| Floatcap | | | -19.993 | 1.930 | | |
| Intercap | | | 8.03 | 0.214 | | |
| Fixedcap | | | -2.928 | 0.902 | | |
| Lenbure | | | | | -7.616 | 2.591 |
| Leninfra | | | | | 5.039 | 1.564 |
| Africa | -0.807 | 1.267 | -0.561 | 1.500 | | |
| Latin America | -0.373 | 0.960 | -0.193 | 0.463 | | |
| Pseudo R squared | 0.157 | | 0.1984 | | 0.225 | |
| No of observations | 674 | | 674 | | 221 | |
| | | Prob | | Prob | | Prob |

| HO: Slopes = 0 | χ^2 (12)= 66.48 | 0.00 | χ^2 (14)=77.42 | $0.00 \gamma^2$ (| 14)=35.01 | 0.00 |
|-----------------------------|----------------------|------|---------------------|--------------------|------------|-------|
| HO: Macro effects 0 | $\chi^2(3) = 13.2$ | 0.00 | χ^2 (3)= 13.34 | | (3)=10.56 | 0.01 |
| HO: Financial effects 0 | χ^2 (4) = 24.81 | 0.00 | χ^2 (4)= 37.14 | ,, , | 4)=13.31 | 0.01 |
| HO: Shock effects $= 0$ | $\chi^2(3) = 19.92$ | 0.00 | χ^2 (7)= 34.29 | | (3) = 6.02 | 0.11 |
| HO: Institu. effects = 0 | 7 | | , , | ,, , | 2) =6.73 | 0.03 |
| Goodness-of-fit: (cut-off p | probability 0.5) | | | Crisis | No crisis | Total |
| Baseline Crisis | | | | 12 | 8 | 20 |
| Predicted crisis | | | | 79 | 575 | 674 |
| Predicted no crisis | | | | 91 | 583 | 674 |
| Total | | | | | | |

6. Sensitivity analysis - different treatment of crisis years

Two variations of crisis years were used in the sensitivity analysis to separate the feedback effects when a crisis was on-going crisis from those factors that influenced the build-up to the crisis. From Table 10, in the first model, all years after the first year of the crisis are deleted. In the second model, all years after the occurrence of the entire crisis are deleted. The results again lend weight to the argument that countries with flexible exchange rates were less likely to face banking crises from external shocks.

Table 10 Definitions of crises

| | All years after 1st crisis year deleted | | All years after entire crises episode deleted | | |
|--|---|----------|---|------------|--|
| Dep. var: Crises | $\delta F(x)/x$ | z | $\delta F(x)/x$ | | |
| Constant | 3.72 | 0.83 | 8.70 | 3.48 | |
| Macro variables | | | | | |
| Rgdpgr | -7.92 | 2.18 | -5.69 | 2.66 | |
| Inflation | 0.22 | 0.37 | 0.15 | 0.33 | |
| Reer | 0.001 | 0.48 | -0.01 | 2.70 | |
| Financial variables | | | | | |
| Lend | -0.09 | 0.06 | 1.00 | 1.26 | |
| Nl2res | 0.02 | 1.01 | 0.02 | 2.10 | |
| Debtgdp | 0.38 | 1.46 | 0.71 | 5.18 | |
| Avintdebt | 0.22 | 2.46 | 0.09 | 1.76 | |
| Shock variables | | | | | |
| Tot trend | | | -2.38 | 4.31 | |
| Floatshock | -0.07 | 2.03 | -0.09 | 3.66 | |
| Intershock | -0.03 | 0.66 | 0.03 | 1.96 | |
| Fixedshock | 0.02 | 1.44 | 0.02 | 2.08 | |
| Floatcap | -1.92 | 0.14 | -20.02 | 2.03 | |
| Intercap | 3.43 | 0.44 | 6.25 | 0.94 | |
| Fixedcap | 4.24 | 0.77 | -3.50 | 1.07 | |
| Psedo R squared | 0.15 | | 0.09 | | |
| No of observations | 540 | | 618 | | |
| | | Prob | | Prob | |
| HO: Slopes = 0 | χ^2 (14)=24.78 | 0.04 | χ^2 (14) =77.75 | 0.00 | |
| HO: Macro effects 0 | $\chi^2(3) = 5.49$ | 0.13 | χ^2 (3)=15.42 | 0.00 | |
| HO: Financial effects 0 | $\chi^2(4) = 9.45$ | 0.05 | χ^2 (4) =35.89 | 0.00 | |
| HO: Shock effects $= 0$ | χ^2 (7)= 11.44 | 0.12 | $\chi^{2}(3)=37.21$ | 0.00 | |
| | , | | | | |
| Goodness of fit models (cut-off probability = 0.5) | | Crisis | No Crisis | Total | |
| Shocks | | | | | |
| Predicted crisis | | 19 | 14 | 22 | |
| Predicted no crisis | | 72 91 | 569 583 | 641 674 | |
| Total | | 91 | 363 | 0/4 | |
| Institutions Crisis | | 17 | 7 | 2.4 | |
| Predicted crisis Predicted no crisis | | 17 28 | 7 169 | 24 197 | |
| Total | | 45 | 176 | 221 | |
| Sensitivity analysis | | | 170 | | |
| Predicted crisis Eq 1 | | 1 | 0 | 1 | |
| Predicted -no crisis | | 21 | 518 | 539 | |
| Total | | 22 | 518 | 540 | |
| Predicted crisis Eq 2 | | | | | |
| Predicted crisis | | 22 | 16 | 38 | |
| Predicted no crisis | | 69 | 511 | 580 | |
| Total | | 91 | 527 | 618 | |

7. Foreign Debt composition

Using the Frankel and Rose [1996] database, the aggregate debt variable was decomposed into long term debt to GDP (Longdebtgdp), commercial debt to GDP, (Comdebtgdp), concessional debt to GDP, (Condebtgdp), variable interest debt to GDP (Vardebt), public debt to GDP, (Pubdebtgdp) and the ratio of short term debt to GDP, (Shodebtgdp). Results are shown in Table 8, starting with the main model.

As seen, only the long term and short term debt definitions were important. They had the expected sign and support the view that excessive levels of debt, especially short term debt, led to problems in the banking sector, while concessional debt was unlikely to have induced crises. These results again corroborate results from the literature, for example by Sachs et al. [1996] and Eichengreen and Rose. The signs on the interacted terms of trade shock variables exhibited the expected signs in the first equation, where more flexible regimes lessened the likelihood of a crisis. The other two equations use different definitions of the crisis periods. Again the results support the underlying story looking at the first and last equation. When all the post-crisis years were deleted in the last equation, the signs on the interacted shock variables remained unchanged. Only the interaction between capital flows and fixed regimes had a sign contrary to what was expected.

Table 11
Debt decomposition

| | Main model | | | All years after Ist crisis | | All years after entire | |
|-------------------------|--------------------------------------|-------|--|----------------------------|----------------------|------------------------|--|
| | | | | year deleted | | crises episode deleted | |
| Dep. var: Crises | $\delta F(x)/x$ | z | $\delta F(x)/x$ | ΙzΙ | $\delta F(x)/x$ | z | |
| Constant | 5.88 | 2.14 | 3.19 | 0.60 | 6.42 | 2.28 | |
| Macro variables | | | | | | | |
| Rgdpgr | -4.58 | 1.99 | -7.89 | 1.83 | -5.30 | 2.24 | |
| Inflation | -1.11 | 1.43 | 0.06 | 0.07 | -0.99 | 1.19 | |
| Reer | 0.00 | 1.06 | 0.00 | 0.06 | 0.00 | 1.23 | |
| Financial variables | | | | | | | |
| Lend | -0.40 | 0.46 | -0.58 | 0.36 | -0.39 | 1.43 | |
| M2res | 0.02 | 1.21 | 0.01 | 0.55 | 0.01 | 0.75 | |
| Londebtgdp | 5.79 | 2.18 | 1.31 | 0.26 | 3.94 | 1.43 | |
| Comdebtgdp | -3.74 | 1.31 | -8.24 | 1.27 | -5.02 | 1.53 | |
| Condebtgdp | -0.72 | -0.50 | 4.92 | 1.60 | 0.21 | 0.14 | |
| Vardebtgdp | 0.47 | 0.16 | 7.83 | 1.23 | 3.34 | 0.97 | |
| Pubdebtgdp | -2.78 | 1.19 | -1.47 | 0.32 | -1.67 | 0.70 | |
| Shodebtgdp | 0.03 | 1.95 | 0.02 | 0.91 | 0.03 | 1.95 | |
| Avintdebt | 0.22 | 3.18 | 0.49 | 3.48 | 0.21 | 2.94 | |
| Shock variables | | | | | | | |
| Tot trend | -2.22 | 3.66 | -2.37 | 1.98 | -2.28 | 3.68 | |
| Floatshock | -0.07 | 2.69 | -0.07 | 1.82 | -0.06 | 2.38 | |
| Intershock | 0.03 | 1.69 | -0.02 | 0.33 | 0.03 | 1.35 | |
| Fixedshock | 0.02 | 1.85 | 0.03 | 1.93 | 0.02 | 2.11 | |
| Floatcap | -17.36 | 1.71 | -10.12 | -0.70 | -15.83 | 1.64 | |
| Intercap | -2.48 | 0.31 | -18.51 | 1.67 | -6.42 | 2.28 | |
| Fixedcap | -5.06 | 1.55 | -1.82 | 0.33 | -6.33 | 0.73 | |
| R squared | 0.27 | | 0.23 | | 0.28 | **** | |
| No of observations | 674 | | 540 | | 618 | | |
| | *** | Prob | | Prob | | Prob | |
| HO: Slopes = 0 | χ^2 (19)= 93.06 | 0.00 | χ^2 (19)=32.16 | 0.03 | χ^2 (19)=91.29 | 0.00 | |
| HO: Macro effects 0 | $\chi^{2}(3) = 7.49$ | 0.06 | $\chi^{2}(3)=3.37$ | 0.03 | $\chi^{2}(3) = 8.33$ | 0.04 | |
| HO: Financial effects 0 | χ^2 (3)= 11.06 | 0.01 | $\chi^{2}(9)=20.20$ | 0.02 | $\chi^{2}(9)=52.95$ | 0.00 | |
| HO: Shock effects = 0 | χ^2 (7)=25.35 | 0.00 | $\chi^2(7) = 10.80$ | 0.15 | $\chi^{2}(3) = 6.02$ | 0.11 | |
| HO: Debt. effects = 0 | $\chi^{2}(6) = 46.36$ | 0.00 | $\chi^{2}(7) = 16.80$ $\chi^{2}(7) = 16.78$ | 0.13 | | 0.03 | |
| 110. Debt. effects – 0 | χ (0) -40.30 | 0.00 | χ (7) - 10.78 | 0.01 | χ^2 (6)=43.58 | 0.03 | |
| | | | | | | | |
| | -(Cut off probability 0 | .5) | | Crisis | No Crisis | Total | |
| Baseline model | Predicted Crisis | | | 30 | 14 | 44 | |
| | Predicted no Crisis | | | 61 | 569 583 | 630 | |
| Shocks model | Total Predicted crisis | | | 91 | 583 0 | 674 2 | |
| Shocks model | Predicted crisis Predicted no crisis | | | 20 | 518 | 538 | |
| | Total | | | 22 | 518 | 540 | |
| Institutions model | Predicted crisis | | | 29 | 14 | 43 | |
| | Predicted no crisis | | | 62 | 513 | 575 | |
| | Total | | | 91 | 527 | 618 | |

8. Fixed effects specification

In the fixed effects version, only those countries which experienced crises were included. The first equation interacted the shocks model with exchange rate regimes. Sensitivity analysis of the crisis variable was carried out following the definitions used previously. The fixed effects results are reported in Table 12 and corroborate some of the findings from the pooled results. The interactive shock terms survives the fixed effects estimation. The capital flow variable is significant but negative in sign. The effects on the financial variables all have robust effects on the fixed effects estimation. However, the significance of the real income growth term diminishes. Of the test for joint significance, only the effects of the macro variables were rejected in the shocks model. Coming to the other two models, all the tests of joint significance are rejected when the first definition of crisis was used, while only the financial effects mattered in the second definition.

Table 12 Fixed effects models

| | Shocks model | | All years after Ist crisis | | All years after entire | |
|-------------------------|----------------------|-------|----------------------------|-------|------------------------|-------|
| | | | year deleted | | crises episode deleted | |
| Dep. var: Crises | $\delta F(x)/x$ | z | $\delta F(x)/x$ | z | $\delta F(x)/x$ | z |
| Macro variables | | | | | | |
| Rgdpgr | -1.96 | 0.613 | -2.445 | 0.144 | 2.330 | 0.570 |
| Inflation | -2.5 | 1.809 | -9.65 | 1.080 | -1.098 | 0.698 |
| Reer | -0.001 | 0.340 | -0.037 | 1.517 | -0.021 | 1.975 |
| Financial variables | | | | | | |
| Lend | 8.215 | 2.692 | 74.247 | 1.930 | 22.074 | 3.169 |
| M2res | 0.056 | 2.911 | 0.299 | 2.088 | 0.164 | 4.214 |
| Debtgdp | 4.905 | 5.275 | 18.433 | 2.529 | 11.657 | 5.038 |
| Avintdebt | 0.374 | 3.547 | 0.597 | 1.686 | 0.321 | 2.107 |
| Shock variables | | | | | | |
| Tot trend | -1.678 | 1.466 | -12.107 | 1.618 | -6.932 | 3.468 |
| Floatshock | -0.051 | 1.702 | 0.093 | 0.803 | 0.024 | 0.497 |
| Intershock | 0.053 | 2.876 | 0.233 | 1.457 | 0.090 | 2.917 |
| Fixedshock | 0.026 | 1.648 | 0.209 | 1.999 | 0.088 | 2.831 |
| Floatcap | -19.118 | 0.901 | -78.849 | 1.065 | -66.122 | 1.763 |
| Intercap | -6.783 | 0.638 | 100.422 | 1.319 | -50.742 | 2.758 |
| Fixedcap | -6.783 | 1.558 | -53.227 | 2.324 | -15.101 | 2.758 |
| R squared | 0.443 | | 0.759 | | 0.71 | |
| No of observations | 387 | | 254 | | 328 | |
| | | Prob | | Prob | | Prob |
| HO: Slopes = 0 | χ^2 (14)= 51.67 | 0.00 | χ^2 (14)=10.72 | 0.70 | χ^2 (14)=33.19 | 0.03 |
| HO: Macro effects 0 | χ^2 (3)=3.47 | 0.32 | χ^2 (3)=2.68 | 0.44 | χ^2 (3)=5.37 | 0.14 |
| HO: Financial effects 0 | $\chi^2(3) = 36.92$ | 0.00 | χ^2 (4)=7.78 | 0.09 | χ^2 (9)=29.13 | 0.00 |
| HO: Shock effects $= 0$ | χ^2 (7)=15.05 | 0.04 | $\chi^2(7) = 6.14$ | 0.52 | χ^2 (3)=16.37 | 0.02 |

V1: Conclusion

This paper examined various determinants of banking crises. The results point to a strong association between the incidence of external shocks and the occurrence of banking crises in SOEs. Key macroeconomic factors such as negative income shocks, level of debt and the real exchange rate were decisive determinants of crises. Countries with high levels of external debt, particularly short term debt were more likely to have banking crises than countries which relied on concessional borrowing. Both terms of trade shocks and capital flows were significant predictors of crises. Some of these factors were also conditioned by the nature of the policy environment in place, in particular the exchange rate regime. This was more profound in cases where external inflows were channeled through fixed or rigid exchange rate regimes. In particular, negative trade shocks, were responsible for a large number of banking crises in the sample. Shocks that were transmitted through more flexible exchange rate regimes caused less problems to the banking sector.

While externally driven factors played a leading role in this process, various other internal disturbances and institutional factors also led to banking crises in SOEs. When low levels of infrastructure and bureaucratic delay were interacted with bank lending, the likelihood of banking crises increased. Again the problem was more acute under rigid exchange rate regimes.

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